

### Section 1.1

- Binomial distribution:  $P(W = w) = \frac{n!}{w!(n-w)!} \pi^w (1-\pi)^{n-w}$  for  $w = 0, 1, \dots, n$  with  $E(W) = n\pi$  and  $\text{Var}(W) = n\pi(1-\pi)$
- $\hat{\pi} = w/n$
- Wald confidence interval for  $\pi$  is  $\hat{\pi} \pm Z_{1-\alpha/2} \sqrt{\frac{\hat{\pi}(1-\hat{\pi})}{n}}$
- Wilson interval for  $\pi$  is  $\tilde{\pi} \pm \frac{Z_{1-\alpha/2} n^{1/2}}{n + Z_{1-\alpha/2}^2} \sqrt{\hat{\pi}(1-\hat{\pi}) + \frac{Z_{1-\alpha/2}^2}{4n}}$  with  $\tilde{\pi} = \frac{w + Z_{1-\alpha/2}^2/2}{n + Z_{1-\alpha/2}^2}$
- Agresti-Coull interval for  $\pi$  is  $\tilde{\pi} \pm Z_{1-\alpha/2} \sqrt{\frac{\tilde{\pi}(1-\tilde{\pi})}{n + Z_{1-\alpha/2}^2}}$
- True confidence level is  $\sum_{w=0}^n I(w) \binom{n}{w} \pi^w (1-\pi)^{n-w}$  where  $I(w) = 1$  if the interval for a  $w$  contains  $\pi$  and 0 otherwise
- Expected length is  $\sum_{w=0}^n L(w) \binom{n}{w} \pi^w (1-\pi)^{n-w}$  where  $L(w)$  is the length of an interval at a particular value of  $w$
- Clopper-Pearson confidence interval for  $\pi$  is  $\text{Beta}(\alpha/2; w, n-w+1) < \pi < \text{Beta}(1-\alpha/2; w+1, n-w)$
- Score test statistic for testing  $\pi = \pi_0$  is  $Z_S = \frac{\hat{\pi} - \pi_0}{\sqrt{\frac{\pi_0(1-\pi_0)}{n}}}$  where a standard normal distribution approximation is used
- General form of LR statistic:  $\Lambda = \frac{\text{Max. lik. when parameters satisfy } H_0}{\text{Max. lik. when parameters satisfy } H_0 \text{ or } H_a}$  and  $-2\log(\Lambda)$  is approximately a  $\chi^2$  random variable for large samples under  $H_0$  with degrees of freedom equal to the difference in dimension between the alternative and null hypothesis parameter spaces

### Section 1.2

- $\hat{\pi}_j = w_j / n_j$
- Wald confidence interval for  $\pi_1 - \pi_2$  is  $\hat{\pi}_1 - \hat{\pi}_2 \pm Z_{1-\alpha/2} \sqrt{\frac{\hat{\pi}_1(1-\hat{\pi}_1)}{n_1} + \frac{\hat{\pi}_2(1-\hat{\pi}_2)}{n_2}}$
- Agresti-Caffo confidence interval is  $\tilde{\pi}_1 - \tilde{\pi}_2 \pm Z_{1-\alpha/2} \sqrt{\frac{\tilde{\pi}_1(1-\tilde{\pi}_1)}{n_1+2} + \frac{\tilde{\pi}_2(1-\tilde{\pi}_2)}{n_2+2}}$  where  $\tilde{\pi}_1 = \frac{w_1+1}{n_1+2}$  and  $\tilde{\pi}_2 = \frac{w_2+1}{n_2+2}$
- Score interval inverts 
$$\frac{|\hat{\pi}_1 - \hat{\pi}_2 - d|}{\sqrt{\hat{\pi}_1^{(0)}(1-\hat{\pi}_1^{(0)})/n_1 + \hat{\pi}_2^{(0)}(1-\hat{\pi}_2^{(0)})/n_2}} < Z_{1-\alpha/2}$$

- Score test statistic for testing  $\pi_1 - \pi_2 = 0$  is  $Z_S = \frac{\hat{\pi}_1 - \hat{\pi}_2}{\sqrt{\bar{\pi}(1-\bar{\pi})\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}$  where a standard normal distribution approximation is used and  $\bar{\pi} = w_+ / n_+$
- Pearson chi-square test statistic for testing  $\pi_1 - \pi_2 = 0$  is  $X^2 = \sum_{j=1}^2 \frac{(w_j - n_j \bar{\pi})^2}{n_j \bar{\pi}(1-\bar{\pi})}$  where a  $\chi^2_1$  distribution approximation is used
- LRT statistic for testing  $\pi_1 - \pi_2 = 0$  is
 
$$-2\log(\Lambda) = -2 \left[ w_1 \log\left(\frac{\bar{\pi}}{\hat{\pi}_1}\right) + (n_1 - w_1) \log\left(\frac{1-\bar{\pi}}{1-\hat{\pi}_1}\right) + w_2 \log\left(\frac{\bar{\pi}}{\hat{\pi}_2}\right) + (n_2 - w_2) \log\left(\frac{1-\bar{\pi}}{1-\hat{\pi}_2}\right) \right]$$
- $RR = \frac{\pi_1}{\pi_2}$  and  $RR = \frac{\hat{\pi}_1}{\hat{\pi}_2} = \frac{w_1 n_2}{w_2 n_1}$
- Wald confidence interval for RR is  $\exp \left[ \log(\hat{\pi}_1 / \hat{\pi}_2) \pm Z_{1-\alpha/2} \sqrt{\text{Var}(\log(\hat{\pi}_1 / \hat{\pi}_2))} \right]$  where
 
$$\text{Var}(\log(\hat{\pi}_1 / \hat{\pi}_2)) = \frac{1}{w_1} - \frac{1}{n_1} + \frac{1}{w_2} - \frac{1}{n_2}$$
- $OR = \frac{\text{odds}_1}{\text{odds}_2} = \frac{\pi_1 / (1-\pi_1)}{\pi_2 / (1-\pi_2)} = \frac{\pi_1(1-\pi_2)}{\pi_2(1-\pi_1)}$  and  $OR = \frac{\text{odds}_1}{\text{odds}_2} = \frac{\hat{\pi}_1(1-\hat{\pi}_2)}{\hat{\pi}_2(1-\hat{\pi}_1)} = \frac{w_1(n_2 - w_2)}{w_2(n_1 - w_1)}$
- Wald confidence interval for OR is  $\exp \left[ \log(OR) \pm Z_{1-\alpha/2} \sqrt{\frac{1}{w_1} + \frac{1}{n_1 - w_1} + \frac{1}{w_2} + \frac{1}{n_2 - w_2}} \right]$
- $OR = \frac{(w_1 + 0.5)(n_2 - w_2 + 0.5)}{(w_2 + 0.5)(n_1 - w_1 + 0.5)}$

## Chapter 2

- Logistic regression model is  $\log\left(\frac{\pi}{1-\pi}\right) = \beta_0 + \beta_1 x_1 + \dots + \beta_p x_p$
- Estimated covariance matrix form when there is only one explanatory variable:
 
$$-E \left[ \begin{pmatrix} \frac{\partial^2 \log[L(\beta_0, \beta_1 | y_1, \dots, y_n)]}{\partial \beta_0^2} & \frac{\partial^2 \log[L(\beta_0, \beta_1 | y_1, \dots, y_n)]}{\partial \beta_0 \partial \beta_1} \\ \frac{\partial^2 \log[L(\beta_0, \beta_1 | y_1, \dots, y_n)]}{\partial \beta_0 \partial \beta_1} & \frac{\partial^2 \log[L(\beta_0, \beta_1 | y_1, \dots, y_n)]}{\partial \beta_1^2} \end{pmatrix} \right]^{-1} \Bigg|_{\beta_0 = \hat{\beta}_0, \beta_1 = \hat{\beta}_1}$$
- Wald test statistic for testing  $\beta_r = 0$  is  $Z_W = \frac{\hat{\beta}_r}{\sqrt{\text{Var}(\hat{\beta}_r)}}$  where a standard normal approximation is used
- LRT statistic for testing a set of  $\beta$ 's are equal to 0 is
 
$$-2\log(\Lambda) = -2\log\left(\frac{L(\beta^{(0)} | y_1, \dots, y_n)}{L(\beta^{(a)} | y_1, \dots, y_n)}\right) = -2 \left[ \sum_{i=1}^n y_i \log\left(\frac{\hat{\pi}_i^{(0)}}{\hat{\pi}_i^{(a)}}\right) + (1 - y_i) \log\left(\frac{1 - \hat{\pi}_i^{(0)}}{1 - \hat{\pi}_i^{(a)}}\right) \right]$$
 where a  $\chi^2_q$  distribution approximation is used (q is number of  $\beta$ 's set to 0 in the null hypothesis)

- $OR = \exp(c\beta_1)$  for an one explanatory variable logistic regression model and a c-unit increase in  $x_1$ ; Wald confidence interval for OR is  $\exp\left(c\hat{\beta}_1 \pm cZ_{1-\alpha/2}\sqrt{\text{Var}(\hat{\beta}_1)}\right)$
- Profile likelihood interval for  $\beta_1$  uses  $-2\log\left(\frac{L(\tilde{\beta}_0, \beta_1 | y_1, \dots, y_n)}{L(\hat{\beta}_0, \hat{\beta}_1 | y_1, \dots, y_n)}\right)$
- In general for two random variables  $W_1$  and  $W_2$  and constants  $a_1$  and  $a_2$ , we have  $\text{Var}(a_1W_1 + a_2W_2) = a_1^2\text{Var}(W_1) + a_2^2\text{Var}(W_2) + 2a_1a_2\text{Cov}(W_1, W_2)$
- Wald confidence interval for  $\pi$  is

$$\frac{\exp\left(\hat{\beta}_0 + \hat{\beta}_1x_1 + \dots + \hat{\beta}_px_p \pm Z_{1-\alpha/2}\sqrt{\text{Var}(\hat{\beta}_0 + \hat{\beta}_1x_1 + \dots + \hat{\beta}_px_p)}\right)}{1 + \exp\left(\hat{\beta}_0 + \hat{\beta}_1x_1 + \dots + \hat{\beta}_px_p \pm Z_{1-\alpha/2}\sqrt{\text{Var}(\hat{\beta}_0 + \hat{\beta}_1x_1 + \dots + \hat{\beta}_px_p)}\right)}$$

where  $\text{Var}(\hat{\beta}_0 + \hat{\beta}_1x_1 + \dots + \hat{\beta}_px_p) = \sum_{i=0}^p \text{Var}(\hat{\beta}_i) + 2\sum_{i=0}^{p-1} \sum_{j=i+1}^p x_ix_j \text{Cov}(\hat{\beta}_i, \hat{\beta}_j)$

- Convergence of the parameter estimates occurs when  $\frac{|G^{(k)} - G^{(k-1)}|}{0.1 + |G^{(k)}|} < \epsilon$  where  $G^{(k)}$  denotes the residual deviance at iteration  $i$
- Probit regression model is  $\text{probit}(\pi) = \beta_0 + \beta_1x_1 + \dots + \beta_px_p$
- Complementary log-log regression model is  $\log[-\log(1 - \pi)] = \beta_0 + \beta_1x_1 + \dots + \beta_px_p$

### Chapter 3

- Multinomial PMF for  $n_1, \dots, n_J$ :  $\frac{n!}{\prod_{j=1}^J n_j!} \prod_{j=1}^J \pi_j^{n_j}$
- $P(X = i, Y = j) = \pi_{ij}$
- Independence:  $\pi_{ij} = \pi_{i+}\pi_{+j}$
- $P(Y = j | X = i) = \pi_{j|i}$
- $\chi^2 = \sum_{i=1}^I \sum_{j=1}^J \frac{(n_{ij} - n_{i+}n_{+j}/n)^2}{n_{i+}n_{+j}/n}$ ; under independence,  $\chi^2$  has a  $\chi^2_{(I-1)(J-1)}$  distribution for a large sample
- $-2\log(\Lambda) = 2\sum_{i=1}^I \sum_{j=1}^J n_{ij} \log\left(\frac{n_{ij}}{n_{i+}n_{+j}/n}\right)$ ; under independence,  $-2\log(\Lambda)$  has a  $\chi^2_{(I-1)(J-1)}$  distribution for a large sample
- P-value for Monte Carlo test involving  $\chi^2$ :  $(\text{Number of } \chi^{2*} \geq \chi^2)/B$
- Multinomial regression model:  $\log(\pi_j/\pi_1) = \beta_{j0} + \beta_{j1}x_1 + \dots + \beta_{jp}x_p$  for  $j = 2, \dots, J$
- For one explanatory variable:  $\pi_1 = \frac{1}{1 + \sum_{j=2}^J \exp(\beta_{j0} + \beta_{j1}x)}$  and  $\pi_j = \frac{\exp(\beta_{j0} + \beta_{j1}x)}{1 + \sum_{j=2}^J \exp(\beta_{j0} + \beta_{j1}x)}$
- Proportional odds model:  $\text{logit}[P(Y \leq j)] = \beta_{j0} + \beta_1x_1 + \dots + \beta_px_p$

- For one explanatory variable:  $\pi_j = \frac{\exp(\beta_{j0} + \beta_1 x)}{1 + \exp(\beta_{j0} + \beta_1 x)} - \frac{\exp(\beta_{j-1,0} + \beta_1 x)}{1 + \exp(\beta_{j-1,0} + \beta_1 x)}$  for  $j = 2, \dots, J - 1$ ,  $\pi_1 = \frac{\exp(\beta_{10} + \beta_1 x)}{1 + \exp(\beta_{10} + \beta_1 x)}$ , and  $\pi_J = 1 - \frac{\exp(\beta_{J-1,0} + \beta_1 x)}{1 + \exp(\beta_{J-1,0} + \beta_1 x)}$
- Nonproportional odds model:  $\text{logit}(P(Y \leq j)) = \beta_{j0} + \beta_{j1}x_1 + \dots + \beta_{jp}x_p$
- Adjacent-categories model:  $\log(\pi_j / \pi_{j+1}) = \beta_{j0} + \beta_{j1}x_1 + \dots + \beta_{jp}x_p$

## Chapter 4

- Poisson PMF:  $P(Y = y) = \frac{\exp(-\mu)\mu^y}{y!}$  where  $E(Y) = \mu$  and  $\text{Var}(Y) = \mu$
- $L(\mu; y_1, \dots, y_n) = \prod_{k=1}^n \frac{\exp(-\mu)\mu^{y_k}}{y_k!}$
- $\hat{\mu} = n^{-1} \sum_{k=1}^n y_k$
- $\text{Var}(\hat{\mu}) = \hat{\mu} / n$
- Score test statistic for testing  $\mu = \mu_0$  is  $Z_s = \frac{\hat{\mu} - \mu_0}{\sqrt{\mu_0 / n}}$  where a standard normal distribution approximation is used
- Score confidence interval for  $\mu$ :  $\left( \hat{\mu} + \frac{Z_{1-\alpha/2}^2}{2n} \right) \pm Z_{1-\alpha/2} \sqrt{\frac{\hat{\mu} + Z_{1-\alpha/2}^2 / 4n}{n}}$
- Wald interval for  $\mu$  using  $\log(\mu)$  transformation:  $\exp(\log(\hat{\mu}) \pm Z_{1-\alpha/2} \sqrt{1/(\hat{\mu}n)})$
- Poisson regression model:  $\log(\mu) = \beta_0 + \beta_1 x_1 + \dots + \beta_p x_p$
- $\mu(x+c) / \mu(x) = \exp(c\beta_1)$  for the model  $\log(\mu) = \beta_0 + \beta_1 x_1$
- $\text{PC} = 100[\exp(c\beta_1) - 1]\%$
- Poisson rate regression model:  $\mu = t \exp(\beta_0 + \beta_1 x)$

## Chapter 5

- $\text{IC}(k) = -2\log(\text{Likelihood function evaluated at parameter estimates}) + kr$
- $\text{AIC} = \text{IC}(2)$
- $\text{AIC}_c = \text{IC}(2n / (n - r - 1)) = \text{AIC} + \frac{2r(r+1)}{n-r-1}$
- $\text{BIC} = \text{IC}(\log(n))$
- $\Delta_m = \text{BIC}_m - \text{BIC}_0$
- $\hat{\tau}_A \approx \frac{\exp(-0.5\Delta_A)}{\sum_{m=1}^M \exp(-0.5\Delta_m)}$
- Model averaged estimate:  $\hat{\theta}_{MA} = \sum_{m=1}^M \hat{\tau}_m \hat{\theta}_m$
- $\text{Var}(\hat{\theta}_{MA}) = \sum_{m=1}^M \hat{\tau}_m [(\hat{\theta}_m - \hat{\theta}_{MA})^2 + \text{Var}(\hat{\theta}_m)]$

- Pearson residual:  $e_m = \frac{y_m - \hat{y}_m}{\sqrt{\text{Var}(Y_m)}}$
- Standardized Pearson residual:  $r_m = \frac{y_m - \hat{y}_m}{\sqrt{\text{Var}(Y_m - \hat{Y}_m)}} = \frac{y_m - \hat{y}_m}{\sqrt{\text{Var}(Y_m)(1 - h_m)}}$
- Deviance residual:  $e_m^D = \text{sign}(y_m - \hat{y}_m)\sqrt{d_m}$
- Standardized deviance residual:  $r_m^D = e_m^D / \sqrt{(1 - h_m)}$
- Pearson statistic:  $X^2 = \sum_{m=1}^M e_m^2 = \sum_{m=1}^M \frac{(y_m - \hat{y}_m)^2}{\text{Var}(y_m)}$
- $D / (M - \tilde{p})$
- Cook's distance:  $CD_m = \frac{r_m^2 h_m}{(p + 1)(1 - h_m)^2}$
- $\Delta X_m^2 = r_m^2$
- $\Delta D_m = (e_m^D)^2 + h_m r_m^2$
- $\text{Var}(Y) = \gamma \mu$
- $\frac{1}{\gamma} \sum_{m=1}^M (Y_m - \mu_m) x_{mr}$
- $r_m = \frac{y_m - \hat{y}_m}{\sqrt{\text{Var}(Y_m)(1 - h_m)\gamma}}$
- $\hat{\gamma} = X^2 / (M - \tilde{p})$
- Negative binomial PMF:  $\binom{y+k-1}{y} \left(\frac{k}{\mu+k}\right)^k \left(1 - \frac{k}{\mu+k}\right)^y$  where  $E(Y) = \mu$  and  $\text{Var}(Y) = \mu + \mu^2/k$

## Chapter 6

- Hypergeometric distribution:  $P(M = m) = \frac{\binom{a}{m} \binom{b}{k-m}}{\binom{n}{k}}$
- Multiple hypergeometric distribution:  $\frac{\left(\prod_{i=1}^I n_{i+}!\right) \left(\prod_{j=1}^J n_{+j}!\right)}{n! \left(\prod_{i=1}^I \prod_{j=1}^J n_{ij}!\right)}$
- P-value for permutation test involving  $X^2$ :  $(\text{Number of } X^{2*} \geq X^2) / B$
- Poisson regression model with random effect:  $\log(\mu_{ik}) = \beta_0 + \beta_1 x_i + b_{0i}$
- $L(\beta_0, \beta_1, \sigma_{b0} \mid y_{11}, \dots, y_{nt}) = \prod_{i=1}^n \int_{-\infty}^{\infty} \prod_{k=1}^t \frac{\exp(-\mu_{ik}) \mu_{ik}^{y_{ik}}}{y_{ik}!} \frac{1}{\sigma_{b0} \sqrt{2\pi}} \exp(-b_{0i}^2 / (2\sigma_{b0}^2)) db_{0i}$
- Predicted value of  $b_{0i}$ :  $E(b_{0i} \mid y_{i1}, \dots, y_{it})$  or mode of  $f(b_{0i} \mid y_{i1}, \dots, y_{it}) = \frac{f(y_{i1}, \dots, y_{it} \mid b_{0i}) \times g(b_{0i})}{f(y_{i1}, \dots, y_{it})}$ ; replace parameters with estimates
- P-value for bootstrap test:  $\frac{1}{B} \sum_{b=1}^B I(W_b^* \geq W)$