

## Equations for STAT 878

1.  $\gamma(s,t) = \text{Cov}(x_s, x_t)$  or  $\gamma(h) = \text{Cov}(x_t, x_{t+h})$
2.  $\rho(s,t) = \frac{\gamma(s,t)}{\sqrt{\gamma(s,s)}\sqrt{\gamma(t,t)}}$  or  $\rho(h) = \frac{\gamma(h)}{\gamma(0)}$
3. For  $x_t = \mu + \sum_{j=-\infty}^{\infty} \psi_j w_{t-j}$  with  $\sum_{j=-\infty}^{\infty} |\psi_j| < \infty$  and  $w_t \sim \text{independent } N(0, \sigma_w^2)$ ,  $\gamma(h) = \sigma_w^2 \sum_{j=-\infty}^{\infty} \psi_{j+h} \psi_j$  for  $h \geq 0$  and  $\gamma(h) = \gamma(-h)$
4.  $\hat{\gamma}(h) = \frac{1}{n} \sum_{t=1}^{n-h} (x_{t+h} - \bar{x})(x_t - \bar{x})$
5.  $\hat{\rho}(h) = \hat{\gamma}(h)/\hat{\gamma}(0)$  has an approximate normal distribution with mean 0 and standard deviation  $\sigma_{\hat{\rho}(h)} = 1/\sqrt{n}$  if  $\rho(h)=0$
6.  $(1 - B)x_t = x_t - x_{t-1}$
7.  $B^k x_t = x_{t-k}$
8.  $\sum_{i=0}^{\infty} a^i = \frac{1}{1-a}$  for  $|a|<1$
9.  $(1-\phi_1 B - \phi_2 B^2 - \dots - \phi_p B^p)(1-B)^d x_t = (1+\theta_1 B + \theta_2 B^2 + \dots + \theta_q B^q)w_t$  and  $\varphi(B)(1-B)^d x_t = \theta(B)w_t$
10.  $\alpha = \mu(1-\phi_1-\phi_2-\dots-\phi_p)$
11. Quadratic formula:  $\frac{-b \pm \sqrt{b^2 - 4ac}}{2a}$  for  $ax^2 + bx + c = 0$
12.  $\gamma(h) = \sum_{j=1}^p \phi_j \gamma(h-j) + \sigma_w^2 \sum_{j=h}^q \theta_j \psi_{j-h}$  for  $0 \leq h < \max(p, q+1)$  and  $\gamma(h) = \sum_{j=1}^p \phi_j \gamma(h-j)$  for  $h \geq \max(p, q+1)$
13.  $\phi_{hh}$  denotes the partial autocorrelation at lag  $h$
14. The estimated partial autocorrelation at lag  $h$ :  
 $\hat{\phi}_{11} = \hat{\rho}(1)$   
 $\hat{\phi}_{hh} = \frac{\hat{\rho}(h) - \sum_{j=1}^{h-1} \hat{\phi}_{h-1,j} \hat{\rho}(h-j)}{1 - \sum_{j=1}^{h-1} \hat{\phi}_{h-1,j} \hat{\rho}(h-j)}$  for  $h = 2, 3, \dots$   
where  $\hat{\phi}_{hj} = \hat{\phi}_{h-1,j} - \hat{\phi}_{hh} \hat{\phi}_{h-1,h-j}$  for  $h=3,4,\dots; j=1,2,\dots,h-1$
15.  $\tilde{x}_{n+m}^n = E(x_{n+m}|I_n)$
16. Forecast error =  $x_{n+m} - \tilde{x}_{n+m}^n$
17.  $\text{Var}[x_{n+m} - \tilde{x}_{n+m}^n] = \sigma_w^2 \sum_{i=0}^{m-1} \psi_i^2$
18.  $(1-\alpha)100\% \text{ C.I. for } x_{n+m}: \tilde{x}_{n+m}^n \pm Z_{1-\alpha/2} \sqrt{\text{Var}(x_{n+m} - \tilde{x}_{n+m}^n)}$
19. Drift term:  $\delta$
20. Standardized residuals:  $e_t = (x_t - \tilde{x}_t^n) / \sqrt{\text{Var}(x_t - \tilde{x}_t^n)}$
21. Ljung-Box-Pierce hypothesis test statistic is  $Q = n(n+2) \sum_{h=1}^H \frac{\hat{\rho}_e^2(h)}{n-h}$ ; under the assumption that the null hypothesis is true,  $Q \sim \chi^2_{H-p-q}$  for large  $n$
22.  $\Phi(B^s)\varphi(B)(1-B^s)^D(1-B)^d x_t = \Theta(B^s)\theta(B)w_t$

$$23. (1-B)^d x_t = \left( \sum_{j=0}^{\infty} \pi_j B^j \right) x_t = \sum_{j=0}^{\infty} \pi_j x_{t-j} \text{ where } \pi_j = \frac{\Gamma(j-d)}{\Gamma(j+1)\Gamma(-d)} \text{ and } \pi_0 = 1$$

$$24. \pi_{j+1} = \frac{(j-d)\pi_j}{(j+1)} \text{ and } \pi_0 = 1$$

$$25. x_t = \sum_{j=0}^{\infty} \psi_j w_{t-j} \text{ where } \psi_j = \frac{\Gamma(j+d)}{\Gamma(j+1)\Gamma(d)}$$

$$26. \rho_{xy}(s,t) = \frac{\gamma_{xy}(s,t)}{\sqrt{\gamma_x(s,s)}\sqrt{\gamma_y(t,t)}}; \text{ if the time series is jointly stationary, then } \rho_{xy}(h) = \frac{\gamma_{xy}(h)}{\sqrt{\gamma_x(0)}\sqrt{\gamma_y(0)}}$$

$$27. \gamma_{xy}(h) = E[(x_{t+h} - \mu_x)(y_t - \mu_y)] = \text{Cov}(x_{t+h}, y_t)$$

$$28. \hat{\gamma}_{xy}(h) = \frac{1}{n} \sum_{t=1}^{n-h} (x_{t+h} - \bar{x})(y_t - \bar{y})$$

29.  $\hat{\rho}_{xy}(h) = \frac{\hat{\gamma}_{xy}(h)}{\sqrt{\hat{\gamma}_x(0)}\sqrt{\hat{\gamma}_y(0)}}$  has an approximate normal distribution with mean 0 and standard deviation

of  $\sigma_{\hat{\rho}_{xy}} = 1/\sqrt{n}$  if the sample size is large and at least one of the series is white noise

$$30. y_t = \beta_0 + \beta_1 z_{t1} + \dots + \beta_r z_{tr} + x_t$$

$$31. y_t = \sigma_t \varepsilon_t \text{ where } \sigma_t^2 = \alpha_0 + \alpha_1 y_{t-1}^2 + \dots + \alpha_m y_{t-m}^2 + \beta_1 \sigma_{t-1}^2 + \dots + \beta_r \sigma_{t-r}^2$$